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Safety and Efficacy of Methotrexate in Psoriasis: A Meta-Analysis of Published Trials

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Abstract

Background

Methotrexate (MTX) has been used to treat psoriasis for over half a century. Even so, clinical data characterising its efficacy and safety are sparse.

Objective

In order to enhance the available evidence, we conducted two meta-analyses, one for efficacy and one for safety outcomes, respectively, according to PRISMA checklist. (Data sources, study criteria, and study synthesis methods are detailed in Methods).

Results

In terms of efficacy, only eleven studies met criteria for study design and passed a Cochrane risk of bias analysis. Based on this limited dataset, 45.2% [95% confidence interval 34.1–60.0] of patients achieve PASI75 at primary endpoint (12 or 16 weeks, respectively, n = 705 patients across all studies), compared to a calculated PASI75 of 4.4 [3.5–5.6] for placebo, yielding a relative risk of 10.2 [95% C.I. 7.1–14.7]. For safety outcomes, we extended the meta-analysis to include studies employing the same dose range of MTX for other chronic inflammatory conditions, e.g. rheumatoid arthritis, in order not to maximise capture of relevant safety data. Based on 2763 patient safety years, adverse events (AEs) were found treatment limiting in 6.9 ± 1.4% (mean ± s.e.) of patients treated for six months, with an adverse effect profile largely in line with that encountered in clinical practice. Finally, in order to facilitate prospective clinical audit and to help generate long-term treatment outcomes under real world conditions, we also developed an easy to use documentation form to be completed by patients without requirement for additional staff time.

Limitations

Meta-analyses for efficacy and safety, respectively, employed non-identical selection criteria.
Conclusions
These meta-analyses summarise currently available evidence on MTX in psoriasis and should be of use to gauge whether local results broadly fall within outcomes.

Introduction
Rationale
Although used for decades to treat psoriasis, surprisingly little systematic review of the use of methotrexate (MTX) in psoriasis has been carried out. A recent meta-analysis, comparing several systemic psoriasis treatments, did include methotrexate [1]. It is instructive that these authors identified approx. 16,000 patients on biologic drugs versus approx. 1800 on conventional systemic drugs, likely reflecting an evidence bias favouring more expensive drugs being subjected to high quality clinical studies. This study reported comparative efficacy of methotrexate, mostly to biologics, while no attempt was made to derive an estimate of PASI75 results versus placebo for methotrexate across clinical trials.

In MTX treatment, safety considerations are often limiting in addition to efficacy. A major safety determinant is the MTX dose employed. If comparable dose ranges are employed, safety outcomes reported in other common chronic inflammatory conditions will thus be informative for psoriasis, even if the pathogenesis of the underlying condition treated differs. We therefore performed a second, extended, meta-analysis to include not only studies on psoriasis but also psoriatic arthritis, rheumatoid arthritis, Crohn’s disease, palmoplantar psoriasis, as well as sero-negative spondylarthropathy. Although the co-morbidity spectrum of these patient populations will be not identical to psoriasis patients, including studies on these related indications in a meta-analysis vastly increases the sample size relevant for safety outcomes.

Pharmaco-economic considerations are increasingly important in guiding treatment decisions. Almost universally, drugs such as MTX will be administered initially in routine practice, only to be replaced by biologics if found ineffective, poorly tolerated, or contra-indicated. Thus, in order to ensure that local practice exhausts all reasonable efforts to achieve sufficient disease control in psoriasis patients before introducing more costly treatments, it is paramount that MTX treatment be audited to establish that local outcomes in terms of safety and efficacy are in line with expected outcomes. However, this presupposes the availability of a point of reference, which currently is lacking.

Objective
We here present meta-analyses of MTX studies, both in terms of efficacy in psoriasis, as well as safety (chronic inflammatory conditions employing same dose range) in order to update currently available evidence on MTX in psoriasis and, at the same time, provide a tentative set of criteria that hopefully will be useful for audit purposes.

Methods
Eligibility criteria
Study selection for the meta-analysis of safety outcomes: Studies selected for safety reporting on methotrexate were identified in accordance with PRISMA guidelines as summarised in
Information sources. An initial search of PubMed, EMBASE, Science Citation Index, and the Cochrane Library was performed using the SEARCH terms in the legend to Fig 1. This set was then reduced by applying the following filter: number of patients on a methotrexate-only treatment arm $\geq 50$, treatment duration $\geq 12$ weeks, dosing oral or intramuscular (i.m.) or subcutaneously (s.c.), RCT design, as well as clearly reported adverse events. This returned 51 entries. Further full-text-based screen resulted in omission of additional 17 studies (IV dosing, non-RCT design, non-reporting of patient numbers on MTX or duration of treatment, patient number below threshold) resulting in a final set of 34 studies (references [2–35]).

Study selection for the meta-analysis of efficacy outcomes: number of patients on methotrexate-only arm $\geq 15$, treatment duration $\geq 12$ weeks, RCT design, and clearly reported psoriasis efficacy outcomes. Of note, the threshold for patient number per trial chosen for efficacy was lower than that for safety ($n = 15$ vs. $n = 50$) purely on pragmatic grounds as application of the more stringent criteria would have left only four studies. For efficacy studies the search terms were: psoriasis/ methotrexate/ trial, yielding an initial set of 45 entries. Filtering of
abstract and title using the criteria above reduced this number to 15. The number was further reduced by two on full-text inspection as several papers reported results from the same trial, did not follow RCT design, had alternative indication (palmoplantar psoriasis), resulting in a final set of thirteen studies (references [3, 23, 25, 36–44]). Three studies (references [3, 23, 25]) fulfil the inclusion criteria for efficacy and safety analysis and are thus included in both datasets.

**Data collection.** All outcomes collected from all individual studies are detailed in **S2** and **S3 Files**.

**Risk of Bias assessment and principle summary measures**

The Cochrane Risk of Bias Assessment tool was used [45] both for the safety studies, as well as efficacy studies. For efficacy outcomes, the risk ratio between MTX and placebo was used (see **Fig 2**).

**Data items and synthesis of results**

Spreadsheets were established to capture both efficacy and safety outcomes (**S4 File**, **S1–S4 Tables**). Incidence figures reported for the various AEs in each study were combined using a random effects model to allow for variability between studies. A random effects (Der-Simonian-Laird) meta-analysis [45] was performed on the incidence rates, yielding a pooled estimate of the mean incidence with 95% confidence interval and standard error. Handling of zeros: an incidence value of 0.5 was added in the case of a zero incidence. In addition, we also explored changing continuity correction of zeros to 0.1 from 0.5 but did not observe any notable difference in the overall results.

For the purpose of analysis AEs were classified into two groups: all AEs vs. treatment limiting AEs. Where an AE was not reported in a study, no imputation of data was made. Patient safety years were calculated for each study using the duration of methotrexate treatment and number of patients available within the studies. The total number of AEs, the total number of severe AEs, the number of AEs likely due to methotrexate treatment and also the number of deaths from each study were also captured. For meta-analysis of PASI75 outcomes in efficacy studies, a similar procedure using an Der-Simonian-Laird estimate was used to allow for calculation of mean AE rate in MTX-treated cohorts relative to patients allocated to placebo arms, followed by calculation of risk ratio between the two pooled estimates (methotrexate and placebo, respectively) as reported in Results.

**Results**

**Methotrexate safety studies**

As detailed above, in order to maximise the available database regarding safety of Methotrexate, the meta-analysis of studies for safety outcomes reported here includes, in addition to psoriasis, indications where a shared dose is used, and with an overlapping co-morbidity and demographic spectrum, that is, psoriatic arthritis, rheumatoid arthritis, Crohn’s disease, palmoplantar psoriasis, as well as sero-negative spondyl-arthropathy. **Table 1** summarises the studies analysed to obtain MTX safety outcomes. The detailed set of all AEs extracted from each study is contained in the Supplement (**S1 Table**). As expected, the number of patients treated for non-psoriasis conditions exceed those for psoriasis and psoriasis arthritis by almost ten-fold. As evident from the table, the dose range used for these indications is comparable to those used in psoriasis, further validating the selection of the above indications for safety analysis.
We identified a total of 34 studies describing methotrexate administration for at least 3 months, applying the search criteria detailed in Methods. Table 2 summarises all AEs reported in at least four studies with an incidence of >0.1%. (The results for all studies are listed in S2 Table.) An unavoidable general limitation of this meta-analysis is that AEs had to be extracted from subcohorts of patients exposed to MTX compared to other treatment arms, e.g. placebo and/or comparator drugs. Thus, the numbers listed do not show excess incidence attributable to MTX exposure, as compared to placebo. The Forest plot of the summary data (Fig 3a) indicates that one factor underlying the high degree of heterogeneity is an apparent under-reporting of AEs in studies with small sample size. Indeed, closer examination of individual AEs reveals that most AE’s appear systematically less frequent in the smaller studies (S4 File).

Methotrexate-associated adverse events (AE)

We identified a total of 34 studies describing methotrexate administration for at least 3 months, applying the search criteria detailed in Methods. Table 2 summarises all AEs reported in at least four studies with an incidence of >0.1%. (The results for all studies are listed in S2 Table.) An unavoidable general limitation of this meta-analysis is that AEs had to be extracted from subcohorts of patients exposed to MTX compared to other treatment arms, e.g. placebo and/or comparator drugs. Thus, the numbers listed do not show excess incidence attributable to MTX exposure, as compared to placebo. The Forest plot of the summary data (Fig 3a) indicates that one factor underlying the high degree of heterogeneity is an apparent under-reporting of AE’s in studies with small sample size. Indeed, closer examination of individual AEs reveals that most AE’s appear systematically less frequent in the smaller studies (S4 File).

Table 1. Published studies reporting MTX safety outcomes relevant for psoriasis treatment.

<table>
<thead>
<tr>
<th>Study</th>
<th>Events</th>
<th>Total</th>
<th>Proportion</th>
<th>95% CI</th>
<th>W(random)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Akhyani 2010</td>
<td>13</td>
<td>18</td>
<td>0.72</td>
<td>[0.47; 0.90]</td>
<td>9.4%</td>
</tr>
<tr>
<td>Barker 2011</td>
<td>90</td>
<td>215</td>
<td>0.42</td>
<td>[0.35; 0.49]</td>
<td>10.1%</td>
</tr>
<tr>
<td>Dogra 2012</td>
<td>18</td>
<td>25</td>
<td>0.72</td>
<td>[0.51; 0.88]</td>
<td>9.7%</td>
</tr>
<tr>
<td>Dogra 2012</td>
<td>24</td>
<td>26</td>
<td>0.92</td>
<td>[0.75; 0.99]</td>
<td>10.2%</td>
</tr>
<tr>
<td>Fallah Arani 2011</td>
<td>6</td>
<td>27</td>
<td>0.22</td>
<td>[0.09; 0.42]</td>
<td>6.3%</td>
</tr>
<tr>
<td>Flystrom 2011</td>
<td>9</td>
<td>37</td>
<td>0.24</td>
<td>[0.12; 0.41]</td>
<td>7.3%</td>
</tr>
<tr>
<td>Goldminz 2015</td>
<td>4</td>
<td>15</td>
<td>0.27</td>
<td>[0.08; 0.55]</td>
<td>5.5%</td>
</tr>
<tr>
<td>Heyendael 2003</td>
<td>26</td>
<td>43</td>
<td>0.60</td>
<td>[0.44; 0.75]</td>
<td>9.7%</td>
</tr>
<tr>
<td>Ho 2010</td>
<td>12</td>
<td>19</td>
<td>0.63</td>
<td>[0.38; 0.84]</td>
<td>9.0%</td>
</tr>
<tr>
<td>Lajevardi 2015</td>
<td>2</td>
<td>22</td>
<td>0.09</td>
<td>[0.01; 0.29]</td>
<td>3.2%</td>
</tr>
<tr>
<td>Reich 2011</td>
<td>59</td>
<td>163</td>
<td>0.36</td>
<td>[0.29; 0.44]</td>
<td>9.9%</td>
</tr>
<tr>
<td>Saurat 2007</td>
<td>40</td>
<td>110</td>
<td>0.36</td>
<td>[0.27; 0.46]</td>
<td>9.7%</td>
</tr>
</tbody>
</table>

Random effects model

Heterogeneity: I-squared=92.7%, tau-squared=0.2015, p<0.0001

0.2 0.4 0.6 0.8

Fig 2. Forest plot of efficacy of MTX in psoriasis as reported in the published clinical studies shown in Table 4. The risk ratio shown refers to the PASI75 outcome reported at 12 or 16 weeks, respectively, assuming a random effects model (http://ije.oxfordjournals.org/content/39/2/421.full). One study (Dogra 2012) is listed twice because two distinct sub-cohorts were dosed differently (10 mg vs 25 mg), as described in the text.

doi:10.1371/journal.pone.0153740.g002

Table 1. Published studies reporting MTX safety outcomes relevant for psoriasis treatment.

<table>
<thead>
<tr>
<th>Indication</th>
<th>All studies(n = 34)</th>
<th>Psoriasis and Pso-Arthritis(n = 5)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Total</td>
<td>Range</td>
</tr>
<tr>
<td>Nr of patients</td>
<td>5995</td>
<td>50–517</td>
</tr>
<tr>
<td>Patient safety years</td>
<td>5083</td>
<td>13–514</td>
</tr>
<tr>
<td>Treatment duration (months)</td>
<td>6 Median</td>
<td>3–24</td>
</tr>
<tr>
<td>Dose (mg)</td>
<td>13.75 Median</td>
<td>7–25</td>
</tr>
</tbody>
</table>

1 Clinical studies were identified and selected as detailed in Methods. Included were studies for psoriasis, psoriasis arthritis, rheumatoid arthritis, and Crohn’s disease. Additional studies for related conditions, e.g., ankylosing spondylitis, were screened (see S1 Table) but removed based on either failed inclusion criteria or risk Cochrane assessment.

2 Range of variables across individual studies.

doi:10.1371/journal.pone.0153740.001
Notably, some—even rare—AE’s, are present with low heterogeneity (S5 File, highlighted in bold print in Table 2). Among these, the incidence of two AEs (pneumonia 0.8%, severe infections, 1.6%), is reported in multiple studies. The frequencies of these AEs therefore appear to be robust against sample size, protocol design, etc and may be most expected to be replicable under real world conditions. The only AE that is highly heterogenous without apparent bias due to study size is nausea/vomiting (S4 File). The reported AE spectrum as such is mostly in line with existing drug labelling information. The most obvious overall limitation in the data is that MTX exposure in published trials rarely exceeds 18 months; more commonly it is limited to 12 months. Therefore, it is impossible to extrapolate the expected incidence of any delayed onset AEs. Another limitation is that reporting of AE terms varies greatly, such that no study reports all of the safety outcomes. Nonetheless, the AE spectrum detailed in Table 2 yields a basis exceeding 1000 patient safety years for the majority of AEs.

### Table 2. Adverse effects associated with MTX treatments in published trials.

<table>
<thead>
<tr>
<th>AE term</th>
<th>Incidence</th>
<th>Range</th>
<th>Studies reporting</th>
<th>Duration (months)</th>
<th>Safety years</th>
</tr>
</thead>
<tbody>
<tr>
<td>All infections</td>
<td>27.6%</td>
<td>3–64.5</td>
<td>14</td>
<td>12 (3–18)</td>
<td>2430</td>
</tr>
<tr>
<td>N/V</td>
<td>18.2%</td>
<td>2–42.6</td>
<td>22</td>
<td>6 (3–24)</td>
<td>3117</td>
</tr>
<tr>
<td>Mouth ulcers</td>
<td>11.1%</td>
<td>0–14</td>
<td>4</td>
<td>12 (3–24)</td>
<td>792</td>
</tr>
<tr>
<td>URI</td>
<td>10.2%</td>
<td>0.6–39</td>
<td>13</td>
<td>11 (6–24)</td>
<td>2424</td>
</tr>
<tr>
<td>Abnormal LFTs</td>
<td>10.0%</td>
<td>1–23.9</td>
<td>17</td>
<td>6 (3–24)</td>
<td>1997</td>
</tr>
<tr>
<td>Abdominal pain</td>
<td>7.5%</td>
<td>1.1–18</td>
<td>8</td>
<td>6 (4–24)</td>
<td>1060</td>
</tr>
<tr>
<td>Headache</td>
<td>7.3%</td>
<td>0.8–27</td>
<td>17</td>
<td>6 (3–24)</td>
<td>2501</td>
</tr>
<tr>
<td>Alopecia</td>
<td>7.3%</td>
<td>2.7–12</td>
<td>5</td>
<td>12 (6–24)</td>
<td>1370</td>
</tr>
<tr>
<td>Diarrhoea</td>
<td>6.8%</td>
<td>1.2–21.6</td>
<td>17</td>
<td>6 (3–24)</td>
<td>2532</td>
</tr>
<tr>
<td>Sinusitis</td>
<td>6.6%</td>
<td>0.2–17</td>
<td>5</td>
<td>12 (6–24)</td>
<td>1471</td>
</tr>
<tr>
<td>Cough</td>
<td>6.4%</td>
<td>2.2–7.5</td>
<td>5</td>
<td>12 (3–24)</td>
<td>877</td>
</tr>
<tr>
<td>Fatigue</td>
<td>6.1%</td>
<td>1.8–16</td>
<td>9</td>
<td>6 (3–18)</td>
<td>895</td>
</tr>
<tr>
<td>Rash</td>
<td>6.0%</td>
<td>0.6–23</td>
<td>8</td>
<td>12 (4–12)</td>
<td>1705</td>
</tr>
<tr>
<td>Dizziness</td>
<td>4.7%</td>
<td>1–11</td>
<td>6</td>
<td>9 (3–24)</td>
<td>854</td>
</tr>
<tr>
<td>Insomnia</td>
<td>4.6%</td>
<td>2.1–5.7</td>
<td>3</td>
<td>4 (3–12)</td>
<td>221</td>
</tr>
<tr>
<td>Leucopenia</td>
<td>3.4%</td>
<td>1–5.9</td>
<td>5</td>
<td>6 (3–18)</td>
<td>265</td>
</tr>
<tr>
<td>UTI</td>
<td>2.9%</td>
<td>0.6–7.2</td>
<td>5</td>
<td>6 (3–11)</td>
<td>537</td>
</tr>
<tr>
<td>Pruritus</td>
<td>2.3%</td>
<td>0–5.6</td>
<td>6</td>
<td>9 (4–24)</td>
<td>1271</td>
</tr>
<tr>
<td>Severe infection</td>
<td>1.6%</td>
<td>0–4.4</td>
<td>14</td>
<td>11.5 (4–24)</td>
<td>2711</td>
</tr>
<tr>
<td>All malignancy</td>
<td>1.2%</td>
<td>0–2</td>
<td>11</td>
<td>12 (5.5–24)</td>
<td>2465</td>
</tr>
<tr>
<td>Pneumonia</td>
<td>0.8%</td>
<td>0–3.9</td>
<td>7</td>
<td>11 (4–24)</td>
<td>1760</td>
</tr>
<tr>
<td>AELTX</td>
<td>28.3%</td>
<td>3.7–52.8</td>
<td>8</td>
<td>6 (5.5–24)</td>
<td>968</td>
</tr>
</tbody>
</table>

1 Data shown include adverse events reported in at least three independent studies with >200 total patient safety years and with a weighted incidence of >0.1%. The full data set containing all reviewed studies is provided in the Supplement. Weighted average rates were calculated using the metaprop procedure in the R package ‘meta’ ((see http://finzi.psych.upenn.edu/library/meta/DESCRIPTION for details). A log transform was used and 0.5 is added to all cell frequencies of studies with a zero cell count.

2 Incidence shown is a weighted incidence to account for the variability of patient numbers across studies, as detailed in Methods.

3 Median duration across studies reporting a given AE.

4 Abbreviations: LFT—liver function tests, N/V—nausea and vomiting; URI—upper respiratory infection; UTI—urinary tract infection.

5 Total patient safety years underlying each AE. The average number of safety years per study is contained the Supplement (S1 Table).

6 AELTX—all adverse events judged as likely due to treatment according to study publication.

doi:10.1371/journal.pone.0153740.t002
Treatment-limiting adverse events reported under MTX exposure

Table 3 details reported treatment-limiting events reported in published trials, which are the ones most relevant to inform clinical practice. Again, heterogeneity is high, as evident from the Forest plots of all individual limiting AE’s (S6 File). However, the systematic trend to under-reporting in small studies is much less pronounced in these more serious AEs, although, as expected for rare events, most limiting AEs are reported in the trials with larger sample sizes.

![Forest plot of all non treatment limiting adverse effects](Fig3a)

![Forest plot of all treatment limiting adverse effects](Fig3b)

All individual study data and forest plots for all individual AE’s are detailed in the Supplement. Underlined studies were conducted for psoriasis as indication.

doi:10.1371/journal.pone.0153740.g003

Table 3. Treatment limiting AEs occurring under methotrexate treatment.

<table>
<thead>
<tr>
<th>AE term</th>
<th>Incidence</th>
<th>Range</th>
<th>Studies reporting</th>
<th>Duration (months)</th>
<th>Safety years</th>
</tr>
</thead>
<tbody>
<tr>
<td>N/V</td>
<td>3.6%</td>
<td>2.2–6.4</td>
<td>2</td>
<td>8 (4–12)</td>
<td>529</td>
</tr>
<tr>
<td>Abnormal LFTs</td>
<td>2.8%</td>
<td>0.6–7.4</td>
<td>7</td>
<td>9 (4–12)</td>
<td>1186</td>
</tr>
<tr>
<td>Leucopenia</td>
<td>2.0%</td>
<td>2–2</td>
<td>2</td>
<td>12 (6–18)</td>
<td>102</td>
</tr>
<tr>
<td>GI</td>
<td>1.1%</td>
<td>0.8–1.7</td>
<td>2</td>
<td>12</td>
<td>680</td>
</tr>
<tr>
<td>Hepatitis</td>
<td>0.8%</td>
<td>0.6–0.9</td>
<td>2</td>
<td>8 (4–12)</td>
<td>200</td>
</tr>
<tr>
<td>Pneumonia</td>
<td>0.8%</td>
<td>0.5–1.1</td>
<td>2</td>
<td>8 (4–12)</td>
<td>213</td>
</tr>
<tr>
<td>Int. pneumonitis</td>
<td>0.7%</td>
<td>0.5–1.0</td>
<td>4</td>
<td>12 (12–24)</td>
<td>1047</td>
</tr>
<tr>
<td>MI</td>
<td>0.6%</td>
<td>0.5–0.7</td>
<td>2</td>
<td>11.5 (11–12)</td>
<td>449</td>
</tr>
<tr>
<td>Alopecia</td>
<td>0.5%</td>
<td>0.2–1.1</td>
<td>2</td>
<td>12</td>
<td>680</td>
</tr>
<tr>
<td>Any Tx lim</td>
<td>6.9%</td>
<td>1.6–28</td>
<td>19</td>
<td>6 (3–24)</td>
<td>2738</td>
</tr>
</tbody>
</table>

1 Data shown include adverse events reported in at least two independent studies with >200 total patient safety years and with a weighted incidence of >0.1%. The full data set containing all reviewed studies is provided in the Supplement. See note 1 for Table 2 above for calculation of weighted averages.

2 Incidence shown is a weighted incidence to account for the variability of patient numbers across studies, as detailed in Methods.

3 Median duration across studies reporting an AE.

4 Abbreviations: N/V—nausea and vomiting; GI—any gastrointestinal event; Int. pneumonitis—interstitial pneumonitis; MI—myocardial infarction.

5 Total patient safety years underlying each AE. The average number of safety years per study is contained the Supplement (S2 Table).

6 Any Tx lim—all treatment limiting adverse events as reported in each study.

doi:10.1371/journal.pone.0153740.t003
(Fig 3b). Of note, the weighted incidences shown are effectively an over-estimate since they can only be calculated based on those studies reporting any such AE, thus leaving out what could well be absence of their occurrence in the remainder of studies. (This is reflected in the greatly reduced overall number of apparent patient-safety years.) Thus, even though the sum of the individual AE incidences, as listed in Table 3, would appear to be 14.6%, the reported incidence of treatment limiting AE’s across all 20 studies reported is 6.9% across all studies (Table 3, bottom row). Although it is likely that studies not reporting a treatment limiting AE in fact didn’t observe it, we cannot make that assumption. Thus, for example, two studies reported treatment-limiting leukopenia whereas the remainder of studies did not. In addition, as with general AE reporting, the influence of dosing and route of administration on treatment-limiting AEs is unclear. Nevertheless, the published data detailed in Table 3 support two conclusions. First, the spectrum of treatment—limiting AE observed in clinical trials (nausea, abnormal LFT, leucopenia, diarrhoea, headache) more or less mirrors that observed in clinical practice. Second, more than 90% of patients receiving MTX in clinical studies do not exhibit any treatment limiting side effects within six months of treatment.

Meta-analysis of efficacy studies on MTX in psoriasis

Table 4 graphically summarises the bias assessment carried out for studies assessed for efficacy assessment. The overall patient numbers as well as the per-study patient numbers are strikingly low compared to studies included for safety variables (937 vs 5995, see above, Table 1). Even this small number of patients mostly consists of patients receiving MTX as comparator treatment for other drugs. For the purpose of efficacy analysis, we did retain partially blinded or unblinded studies (Table 4, marked by ‘x’). The rationale behind this is that, in contrast to the analysis reported in [1], we did not aim at defining comparative efficacy of MTX versus other treatments but rather at defining an efficacy outcome informative for routine practice under real world conditions. Therefore, any potential bias in favour of MTX introduced by including incompletely blinded studies is more than likely offset by the inherent bias against MTX by omission of two key variables used in routine care to optimise response to MTX, namely individual dose adjustment as well as alteration of route of administration. Two studies found to have significant issues were omitted from further analysis (Table 4). Taken together, the published database allowing any assessment of MTX efficacy in psoriasis is strikingly small despite decades of clinical use.

Efficacy outcomes of MTX in psoriasis

Despite the obvious limitations of limited study numbers, divergent study design, non-uniform outcome reporting, and small patient numbers, we performed a meta-analysis of treatment efficacy based on the only variable that was accessible across all studies analysed: the percentage of patients achieving 75% reduction of PASI from baseline (PASI75). Fig 2 graphically summarises the PASI75 reported in the MTX-only treatment arm at 12 or 16 weeks, respectively, in each of the studies analysed. As evident from the figure, there is notable heterogeneity between studies ($I^2 = 92.7\%$). Dose range as underlying factor is unlikely, given that many studies employed a flexible dose increase scheme, covering a brought overall range (for details, see S4 File). We were also unable to identify any other systematic potential causes, e.g. sample size, dosing scheme. Therefore, in the absence of any discernable factor accounting for heterogeneity, the figure displays a random effects model. The pooled PASI75 estimate calculated across all studies yielded at PASI75 of 45.2% (95% confidence interval 34.1%–60.0%) compared to a calculated PASI75 of 4.4% for placebo (95% confidence interval 3.5%–5.6%) (S4 Table). This yields a relative risk of 10.2 (95% confidence interval 7.1–14.7). Bearing in mind the caveats
detailed above, we conclude that, across published MTX studies on psoriasis, approximately 40% of patients achieve PASI75 between weeks 12–16.

**Discussion**

**Summary of evidence**

Few would argue that methotrexate is as effective as modern biologics for the treatment of moderate to severe psoriasis. Its comparative inferiority has previously been summarised in an in-depth meta-analysis [1]. However, when trying to determine just how effective it is, or just how common side effects occur, we are confronted with a limited database. In terms of safety, the present analysis uncovers a systematic bias toward under-reporting of adverse effects in studies of small sample size (S4 File). This observation adds support to our inclusion of studies using the same dose range as used for psoriasis in similar indications (i.e. chronic inflammatory, non-cancer indications). If one were to restrict safety analysis to studies on psoriasis, many of the adverse effects clinically relevant for psoriasis patients simply would not be detectable due to limiting cohort sizes. Moreover, one inherent overall limitation is the duration of published studies. In general, these are too short to detect delayed-onset adverse effects. Thus, only real-world data would be able to identify frequencies of late-onset adverse effects occurring with long-term dosing.

The overall spectrum of adverse effects mirrors that observed in routine practice. Regarding treatment limiting AE’s it is worth noting that, despite common perception, the most common limiting effects, LFT aberration and nausea, are rare. The latter was only reported in 2 / 20 trials, somewhat reducing the statistical significance of the actual incidence. Limiting LFT aberration also occurs infrequently (3.1% of patients, reported in 8 out of 22 studies). These findings are in agreement with a meta-analysis of liver-injury in methotrexate users, which found that patients on methotrexate are not at increased risk of liver cirrhosis or liver failure [46].

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**Table 4. Assessment of bias-profile in published trials on MTX efficacy in psoriasis.**

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¹ Clinical trials reporting efficacy data of MTX in psoriasis were identified by literature search as detailed in Methods. Studies are as follows: A Akhyani 2010, B Barker 2011, C Dogra 2012, D Fallah Arani 2011, E Flystrom 2007, F Goldminz 2015, G Heyendael 2003, H Ho 2010, I Lajevardi 2015, J Radmenesh 2011, K Ranjan 2007, L Reich 2011, M Saurat 2007 Bias-profile was analysed using the Cochrane Risk of Bias tool (see Methods). Coding is: 1—low risk, 2—high risk, 3—unclear risk. Studies Radmenesh 2011 and Ranjan 2007 were omitted from further analysis based on unacceptably high bias profile.

² Shown are the number of patients reported in each study allocated to the methotrexate—only arm.

³ This study reports two separate sub-cohorts at different dosing levels, respectively.

⁴ DLQI—Dermatology Life Quality Index [51]
In terms of efficacy, the present meta-analysis finds that approximately 40% of patients achieve a PASI75 by 12 weeks of treatment. As with adverse effect, the main issue evident from the data is the marked heterogeneity, the spread ranging between 9% and 92% of patients reaching this threshold (Fig 2). One factor underlying this heterogeneity would appear to be study design, for example, single-blinded design (see Table 3). Other possible causes are ethnic heterogeneity between patient cohorts, precise dosing scheme, or confounding metabolic factors. Thus, pre-treatment calcium levels were most recently found to be a predictor of treatment response [47]. Clearly, this number is much lower than that achieved by modern biologics, as detailed in [1]. Nonetheless, it means that many psoriasis patients requiring systemic therapy are likely to benefit from methotrexate, even more so, as flexible dose adjustment and adaptation of route of dosing are applied under real-world conditions which would be expected to increase efficacy and reduce limiting adverse events.

Most recently, both FDA and EMA granted a licence for the biologic secukinumab for psoriasis. Significantly, the licence was granted as first-line systemic treatment (www.fda.gov/downloads/drugs/drugsafety/ucm433352.pdf), in principle enabling clinicians to choose this costly drug instead of the widely used first-line low cost option of MTX. In light of development such as this one, it will become increasingly important for health care providers to understand if reasonable efforts are being made to employ cost-effective treatments wherever safe and effective. To this end, benchmarks need to be developed under real world conditions in order to allow audit of treatment outcomes. To assist with this, we developed a documentation tool (S7 File). Since staff time is limiting in most routine clinical settings, the tool has been designed to be completed by patients, with clinician input limited to only recording of global psoriasis state (PGA) and LFT aberration. In terms of safety outcomes, the tool exploits the fact that most adverse effect in fact represent patient reported outcomes (see Tables 2 and 3). In terms of efficacy, since completion of PASI scores is rarely feasible in routine clinics, we selected the much easier to use commonly employed PGA scale. This is informative, since PASI75 closely approximates PGA 0/1 in clinical trial settings [48] and since PASI and PGA exhibit close correlation across all severity bands [49]. The tool is easily adaptable to electronic documentation systems wherever these are used (e.g., using Google-Forms). Aside from audit, this tool should also allow prospective collection of data with minimal selection and reporting bias.

Limitations
One important limitation of the present report is that, in terms of safety profile, we present two meta-analyses, where the analysis of safety includes indications other than psoriasis. The rationale for this decision is that (i) the primary safety determinant of methotrexate is the dose range employed, which is identical between the non-psoriasis and psoriasis studies and (ii) that no published data have been reported to indicate that the pharmaco-responsiveness between psoriasis and non-psoriasis patients deviates in the general population for any drug, thereby rendering safety data from non-psoriasis trials relevant for psoriasis patients. However, we cannot exclude the theoretical possibility that, with sufficiently large data to hand, the safety profile between psoriasis and non-psoriasis cohorts would diverge.

Conclusions
The present meta-analysis vividly illustrates and reinforces the limitations of clinical trials as basis of policy and guideline drafting. Compared to biologics, only minimal evidence is available on treatment efficacy of methotrexate in psoriasis. Since this is unlikely to change, given funding constraints for clinical trials, real-world data, although inherently fuzzy, incomplete,
and inevitably less well documented, are the sole source for understanding the performance of drugs such as methotrexate in their intended use, e.g. longterm [50].

Supporting Information
S1 File. PRISMA Checklist.
(PDF)
S2 File. Detailed-outcomes-SAFETY.
(XLSX)
S3 File. Detailed-outcomes-EFFICACY.
(XLSX)
S4 File. Forest plots for non-treatment limiting AE’s with high heterogeneity.
(TIF)
S5 File. Forest plots for non-treatment limiting AE’s with low heterogeneity.
(TIF)
S6 File. Forest plots for treatment limiting AE’s.
(TIF)
S7 File. MTX Treatment Documentation Tool.
(DOCX)
S1 Table. Adverse effects associated with MTX treatments in published trials.
(DOCX)
S2 Table. Treatment limiting AE’s occurring under methotrexate treatment.
(DOCX)
S3 Table. Summary safety outcome statistics from published data.
(DOCX)
S4 Table. Data on methotrexate placebo arm—allocated patients.
(DOCX)

Author Contributions
Conceived and designed the experiments: JF JW. Performed the experiments: JW SO. Analyzed the data: JW SO JF. Wrote the paper: JF SO JW.

References


